

# Exponential Stability of Modified Stochastic Approximation Procedure <sup>\*†</sup>

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Received 2 February 2002

## Abstract

The modified stochastic approximation procedure

$$x_{i+1} = x_i + \alpha_i g(x_i + \xi_{i+1}), x_0 = \zeta \quad (1)$$

is considered. Here  $\{\alpha_j\}$  is the sequence of positive numbers,  $\{\xi_n\}$  is a sequence of martingale-differences, function  $g$  is twice differentiable,  $ug(u) \leq 0$  for  $u \neq 0$ ,  $\zeta$  is the initial value. Results on the almost-sure boundedness and the exponential stability of procedure (1) are obtained. The theorem of the convergence of nonnegative semimartingale has been applied.

## 1 Introduction

Stochastic approximation, originally proposed by Robbins and Monro in 1951 [7], is concerned with the problem of finding the root of the function  $y = R(x)$  which is neither known nor directly observed. Let the result of the measurement at the point  $x_k$  at moment  $k$  be equal to  $Y_k = R(x_k) + \xi_{k+1}$ , where  $\xi_1, \dots, \xi_k, \dots$  are independent random values with zero mean. For an arbitrary initial point  $X(0) = x$  and an arbitrary sequence  $\{\gamma_k\}$  of positive numbers, Robbins and Monro suggested the following procedure

$$X_{k+1} = X_k - \gamma_k Y_k, Y_k = R(x_k) + \xi_{k+1}. \quad (2)$$

The generalizations of Robbins and Monro method were investigated in numerous publications. We mention here just two outstanding books: by Nevelson and Khasminskii (cf. [5]) and Kushner (cf. [2]). Over the years, stochastic approximation has been proven to be a powerful and useful tool. Here we discuss the application of the stochastic approximation to credibility.

Let  $x$  denote the claim size, with distribution  $P_\theta$  that depends on some random parameter  $\theta$  (with the prior distribution  $\pi(\theta)$ ). It can be proven (cf. [1] and [3]) that

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\*Mathematics Subject Classifications: 60H10, 93E15.

†Partially supported by Cariscience

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for some distributions  $P$  and  $\pi$  (the Normal/Normal, the Poisson/Gamma, *etc.*) the tradition credibility formula

$$\hat{\mu}_n = (1 - \alpha_n)m + \alpha_n \bar{x}_n, \quad (3)$$

takes place. Here  $\hat{\mu}_n$  is an estimation of the fair premium,  $m$  is the collective fair premium,  $\bar{x}_n$  is the mean of  $n$  years of individual experience  $x_1, x_2, \dots, x_n$ . It is not difficult to show that (3) can be rewritten as a stochastic recursion of the type (2). The latter is particularly suited for the sequential evaluation of the fair premium. However in some situations (see [3]) tradition credibility formula fails and the stochastic approximation gives rise to some kind of quasi-credibility. In this case, instead of the stochastic approximation procedure (2), we need to consider the following modified procedure

$$x_{i+1} = x_i + \alpha_i g(x_i + \xi_{i+1}), \quad x_0 = \zeta. \quad (4)$$

In this paper we investigate the exponential stability of procedure (4), where the errors of the observation  $\xi_i$  are martingale-differences, the function  $g$  is twice differentiable and  $ug(u) \leq 0$  for  $u \neq 0$ . The theorem of the convergence of nonnegative semimartingale will be needed (cf. [6]).

## 2 Definitions and Auxiliary Lemmas

Let the probability space  $(\Omega, F, P)$  with filtration  $F = \{\mathcal{F}_n\}_{n=1,2,\dots}$  be given. Let the stochastic sequence  $\{m_n\}$  be an  $\mathcal{F}_n$ -martingale with  $m_0 = 0$ . We put  $\xi_n = m_n - m_{n-1}$  for  $n \geq 1$  where  $\xi_0 = 0$ . Then the stochastic sequence  $\{\xi_n\}$  is an  $\mathcal{F}_n$ -martingale-difference. For detailed definitions and facts of random processes, the reader can see e.g. [4]. We present below two necessary lemmas which will be used in this paper.

LEMMA 1. Let  $\{\xi_n\}$  be an  $\mathcal{F}_n$ -martingale-difference. Then there exists an  $\mathcal{F}_n$ -martingale-difference  $\{\mu_n\}$  and a positive  $\mathcal{F}_{n-1}$ -measurable stochastic sequence  $\{\eta_n\}$  such that for every  $n = 1, 2, \dots$ ,

$$\xi_n^2 = \mu_n + \eta_n \quad \text{a.s.} \quad (5)$$

LEMMA 2. Let  $Z_n = Z_0 + A_n^1 - A_n^2 + M_n$  be a non-negative semimartingale, where  $M_n$  is a martingale,  $A_n^1, A_n^2$  are a.s. non-decreasing processes. Then  $\{\omega :: A_\infty^1 < \infty\} \subseteq \{Z \rightarrow\} \cap \{A_\infty^2 < \infty\}$  a.s.

Here  $\{Z \rightarrow\}$  denotes the set of all  $\omega \in \Omega$  for which  $Z_\infty = \lim_{t \rightarrow \infty} Z_t$  exists and is finite. The notation a.s. means almost-surely.

## 3 Boundedness of Solution

Let the function  $g$  be twice differentiable and for any  $u \in \mathfrak{R}$ ,

$$ug(u) \leq 0, \quad u \neq 0, \quad (6)$$

$$g^2(u) \leq K_1 u^2, \quad (7)$$

$$|g''(u)| \leq K, \quad (8)$$

where  $K_1, K > 0$  are nonrandom numbers. Let  $\{\xi_n\}$  be a sequence of  $\mathcal{F}_n$ -martingale-differences with  $\xi_0 = 0$  and decomposition (5) takes place. Let  $\{\alpha_j\}$  be a sequence of positive numbers such that a.s.

$$\sum_{j=0}^{\infty} \alpha_j = \infty, \quad (9)$$

$$\sum_{j=0}^{\infty} \alpha_j^2 < \infty, \quad (10)$$

$$\sum_{j=0}^{\infty} \alpha_j \eta_{j+1} < \infty. \quad (11)$$

**THEOREM 1.** Let conditions (6)-(11) be fulfilled, then the solution  $x_i$  to equation (4) has the following properties:

$$\sup_{0 < i < \infty} x_i^2 < \infty \quad \text{and} \quad \liminf_{i \rightarrow \infty} x_i = 0 \quad \text{a.s.}$$

**PROOF.** Applying Taylor's expansion to  $g(x_i + \xi_{i+1})$  we have

$$\begin{aligned} x_{i+1}^2 - x_i^2 &= (x_i + \alpha_i g(x_i + \xi_{i+1}))^2 - x_i^2 \\ &\leq 2\alpha_i x_i \left[ g(x_i) + g'(x_i)\xi_{i+1} + g''(u)\frac{\xi_{i+1}^2}{2} \right] + \alpha_i^2 K_1 (x_i + \xi_{i+1})^2, \end{aligned}$$

where  $u$  lies between  $x_i$  and  $\xi_{i+1}$ . From above and using the decomposition of  $\xi_i^2$  (see Lemma 1 and (5)) we have

$$\begin{aligned} x_{i+1}^2 - x_i^2 &\leq 2\alpha_i x_i g(x_i) + (K\alpha_i |x_i| + 2K_1 \alpha_i^2) \eta_{i+1} + 2K_1 \alpha_i^2 x_i^2 + 2\alpha_i x_i g'(x_i) \xi_{i+1} \\ &\quad + (K\alpha_i |x_i| + 2K_1 \alpha_i^2) \mu_{i+1}. \end{aligned}$$

Let  $\Delta m_i = 2\alpha_i x_i g'(x_i) \xi_{i+1} + (K\alpha_i |x_i| + 2K_1 \alpha_i^2) \mu_{i+1}$ , which is a martingale-difference. Applying the estimation  $|x_i| \leq 1 + x_i^2$ , we get

$$x_{i+1}^2 - x_i^2 \leq 2\alpha_i x_i g(x_i) + (K\alpha_i + 2K_1 \alpha_i^2) \eta_{i+1} + (K\alpha_i \eta_{i+1} + 2K_1 \alpha_i^2) x_i^2 + \Delta m_i, \quad (12)$$

and then

$$x_{i+1}^2 \leq 2\alpha_i x_i g(x_i) + (1 + \beta_i) x_i^2 + (K\alpha_i + 2K_1 \alpha_i^2) \eta_{i+1} + \Delta m_i, \quad (13)$$

where  $\beta_i = K\alpha_i \eta_{i+1} + 2K_1 \alpha_i^2$ . Note that (10)-(11) imply

$$\prod_{j=1}^i (1 + \beta_j) < M \quad (14)$$

for some nonrandom  $M > 0$  and every  $i = 1, 2, \dots$ . Letting

$$x_{i+1} = \prod_{j=1}^i (1 + \beta_j)^{1/2} y_{i+1}$$

and substituting it in (13) we get

$$\begin{aligned} \prod_{j=1}^i (1 + \beta_j)(y_{i+1}^2 - y_i^2) &\leq 2\alpha_i \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i g \left( \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i \right) \\ &\quad + (K\alpha_i + 2K_1\alpha_i^2)\eta_{i+1} + \Delta m_i. \end{aligned}$$

Let

$$\prod_{j=1}^i (1 + \beta_j)^{-1} \Delta m_i = \Delta m_i^1,$$

which is a martingale-difference, therefore

$$\begin{aligned} y_{i+1}^2 - y_i^2 &\leq 2 \prod_{j=1}^i (1 + \beta_j)^{-1} \alpha_i \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i g \left( \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i \right) \\ &\quad + \prod_{j=1}^i (1 + \beta_j)^{-1} (K\alpha_i + 2K_1\alpha_i^2)\eta_{i+1} + \Delta m_i^1. \end{aligned} \quad (15)$$

Taking the sum of (15) from  $i = 0$  to  $i = n - 1$ , we have

$$\begin{aligned} \sum_{i=0}^{n-1} y_{i+1}^2 - \sum_{i=0}^{n-1} y_i^2 &\leq 2 \sum_{i=0}^{n-1} \prod_{j=1}^i (1 + \beta_j)^{-1} \alpha_i \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i g \left( \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i \right) \\ &\quad + \sum_{i=0}^{n-1} \prod_{j=1}^i (1 + \beta_j)^{-1} (K\alpha_i + 2K_1\alpha_i^2)\eta_{i+1} + \sum_{i=0}^{n-1} \Delta m_i^1. \end{aligned}$$

Therefore

$$y_n^2 \leq U_n = y_0^2 + A_n^1 - A_n^2 + m_n^1, \quad (16)$$

where

$$A_n^1 = \sum_{i=0}^{n-1} \prod_{j=1}^i (1 + \beta_j)^{-1} (K\alpha_i + 2K_1\alpha_i^2)\eta_{i+1}$$

and

$$A_n^2 = -2 \sum_{i=0}^{n-1} \prod_{j=1}^i (1 + \beta_j)^{-1} \alpha_i \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i g \left( \prod_{j=1}^{i-1} (1 + \beta_j)^{1/2} y_i \right).$$

It should be noted that  $A_n^1$  and  $A_n^2$  are increasing processes a.s.,  $m_n^1$  is a martingale,  $U_n$  is nonnegative semimartingale and  $P\{A_\infty^1 < \infty\} = 1$  due to the convergence of series in (10) and (11). Applying Lemma 2, that is

$$\{A_\infty^1 < \infty\} \subset \{U_i \rightarrow\} \cap \{A_\infty^2 < \infty\},$$

we have

$$P\{U_i \rightarrow\} = 1.$$

This implies that there exists some a.s. finite random value  $H = H(\omega)$  such that  $P\{\sup_{0 < i < \infty} U_i \leq H\} = 1$ . Then  $P\{\sup_{0 < i < \infty} y_i^2 \leq H\} = 1$  and the first part of the theorem is proved.

Suppose now that  $P\{\liminf_{i \rightarrow \infty} y_i^2 > 0\} = p_0 > 0$ . Then there exist random variables  $\zeta_0 = \zeta_0(\omega) > 0$  and  $N_0 = N_0(\omega) > 0$  such that  $P(\Omega_0) = p_0$ , where  $\Omega_0 = \{\omega : y_i^2 > \zeta_0(\omega)/2 \text{ for } i > N_0\}$ . Since  $\prod_{j=1}^i (1 + \beta_j)^{1/2} > 1$ , we have  $y_i^2 \prod_{j=1}^i (1 + \beta_j)^{1/2} > \zeta_0(\omega)/2$  for  $i > N_0$  and  $\omega \in \Omega_0$ . Due to the continuity and negativity (for  $u \neq 0$ ) of the function  $\phi(u) = ug(u)$  we can find  $k_0 = k_0(\omega)$  and  $N_1 = N_1(\omega) \geq N_0(\omega)$  such that

$$-\phi\left(\prod_{j=1}^i (1 + \beta_j)^{1/2} y_i\right) > k_0(\omega)$$

for  $\omega \in \Omega_0$  and  $i > N_1$ . Then

$$\begin{aligned} & -2 \sum_{i=0}^{n-1} \alpha_i \prod_{j=1}^i (1 + \beta_j)^{1/2} y_i(\omega) g\left(\prod_{j=1}^i (1 + \beta_j)^{1/2} y_i(\omega)\right) \\ = & -2 \sum_{i=0}^{N_1-1} -2 \sum_{i=N_1}^{n-1} \\ \geq & -2 \sum_{i=N_1}^{n-1} \alpha_i \phi\left(\prod_{j=1}^i (1 + \beta_j)^{1/2} y_i(\omega)\right) \\ \geq & 2k_0(\omega) \sum_{i=N_1}^{n-1} \alpha_i \rightarrow \infty \end{aligned}$$

as  $n \rightarrow \infty$ . Hence  $P\{A_\infty^2 = \infty\} \geq p_0 > 0$  which contradicts (16) and conditions (10)-(11). Then  $P\{\liminf_{i \rightarrow \infty} y_i^2 > 0\} = 0$  and from (14) we have:  $P\{\liminf_{i \rightarrow \infty} x_i^2 \leq M \liminf_{i \rightarrow \infty} y_i^2 = 0\} = 1$ . The theorem is completely proved.

## 4 Exponential Stability

In addition to the conditions from the previous section let two more conditions be fulfilled

$$H_1 |x| \leq |g(x)|, \quad (17)$$

$$\sum_{i=0}^{\infty} \alpha_i \eta_{i+1} \prod_{j=1}^i (1 - 2\alpha_j H_1)^{-1} < \infty \text{ a.s.}, \quad (18)$$

where  $H_1 > 0$  is some constant and  $\eta_{i+1} \rightarrow 0$  when  $i \rightarrow \infty$ .

REMARK 1. Let  $H_1 |x| \leq |g(x)|$  and  $xg(x) < 0$  for all  $x \in \mathfrak{R}$  and  $x \neq 0$ . The following is correct: if  $x > 0$ , then

$$-xg(x) = x|g(x)| \geq |x|H_1|x| = H_1 x^2;$$

and if  $x < 0$ , then

$$-xg(x) = |x||g(x)| \geq |x|H_1|x| = H_1x^2.$$

Therefore

$$xg(x) \leq -H_1x^2. \quad (19)$$

**THEOREM 2.** Let conditions (6)-(11) and (17)-(18) be fulfilled. Then for any  $\kappa > 0$ ,

$$\exp \left[ (1 - \kappa)2H_1 \sum_{i=0}^n \alpha_i \right] x_n^2 \rightarrow 0$$

a.s. where  $x_n$  is a solution of equation (4).

**PROOF.** Substituting (19) in (12) we get

$$x_{i+1}^2 - x_i^2 \leq -(2\alpha_i H_1 - K\alpha_i \eta_{i+1} - 2K_1 \alpha_i^2) x_i^2 + (K\alpha_i + 2K_1 \alpha_i^2) \eta_{i+1} + \Delta m_i. \quad (20)$$

Let

$$\tau_i = 2\alpha_i H_1 - K\alpha_i \eta_{i+1} - 2K_1 \alpha_i^2.$$

Due to the positivity of  $K_1$ ,  $K$  and  $\eta_{i+1}$  from condition (18) we have

$$\begin{aligned} & \sum_{i=0}^{\infty} (K\alpha_i - 2K_1 \alpha_i^2) \eta_{i+1} \prod_{j=0}^i (1 - 2\alpha_j H_1 + 2K_1 \alpha_j^2 + K\alpha_j \eta_{j+1})^{-1} \\ & \leq K \sum_{i=0}^{\infty} \alpha_i \eta_{i+1} \prod_{j=0}^i (1 - 2\alpha_j H_1)^{-1} < \infty. \end{aligned} \quad (21)$$

From (20) we get

$$x_{i+1}^2 \leq (1 - \tau_i) x_i^2 + (K\alpha_i + 2K_1 \alpha_i^2) \eta_{i+1} + \Delta m_i. \quad (22)$$

Let

$$z_i^2 = \prod_{j=1}^{i-1} (1 - \tau_j)^{-1} x_i^2$$

which implies that

$$x_i^2 = \prod_{j=1}^{i-1} (1 - \tau_j) z_i^2.$$

Substituting it in (22) we get,

$$\prod_{j=1}^i (1 - \tau_j) z_{i+1}^2 \leq (1 - \tau_i) \prod_{j=1}^{i-1} (1 - \tau_j) z_i^2 + (K\alpha_j + 2K_1 \alpha_j^2) \eta_{j+1} + \Delta m_i$$

and

$$z_{i+1}^2 - z_i^2 \leq \prod_{j=1}^i (1 - \tau_j)^{-1} (K\alpha_j + 2K_1 \alpha_j^2) \eta_{j+1} + \Delta m_i^1, \quad (23)$$

where  $\Delta m_i^1 = \prod_{j=1}^i (1 - \tau_j)^{-1} \Delta m_i$ . Taking the sum of (23) from  $i = 0$  to  $i = n - 1$ , we obtain

$$z_n^2 \leq z_0^2 + \sum_{i=0}^{n-1} \prod_{j=1}^i (1 - \tau_j)^{-1} (K\alpha_j + 2K_1\alpha_j^2)\eta_{j+1} + m_n^1.$$

Applying (21) and using the same arguments as in Theorem 1 we obtain that there exists some a.s. finite random value  $H = H(\omega)$  such that  $P\{\sup_{0 < i < \infty} z_i^2 \leq H\} = 1$ . As  $\alpha_i, \eta_{i+1} \rightarrow 0$  when  $i \rightarrow \infty$  for any  $\varepsilon > 0$  there exists the random integer  $N = N(\omega) > 0$  such that for any  $j \geq N$

$$2K_1\alpha_j + K\eta_{j+1} < 2H_1\varepsilon.$$

Then for some  $H_2 > 0$

$$\begin{aligned} x_i^2 &\leq HH_2 \prod_{j=N}^i (1 - 2H_1(1 - \varepsilon)\alpha_j) = HH_2 \exp\left\{\sum_{j=N}^i \ln[1 - 2H_1(1 - \varepsilon)\alpha_j]\right\} \\ &\leq HH_2 \exp\left\{-2H_1 \sum_{j=N}^i (1 - \varepsilon)\alpha_j\right\}. \end{aligned} \quad (24)$$

If we take some  $\kappa > 0$  and  $\varepsilon < \kappa/2$ , then from (24) we obtain that  $P\{\lim_{i \rightarrow \infty} \exp\{(1 - \kappa)2H_1 \sum_{j=N}^i \alpha_j\} x_i^2 = 0\} = 1$ . The proof is complete.

In the following example we investigate the fulfillment of the condition (18). We consider two different cases for  $\alpha_i$ , and  $\eta_i$ .

EXAMPLE. a) Let  $\alpha_i = \frac{1}{i}$  and  $\eta_i \leq C(\omega)/i^{\varepsilon+2H_1}$  for some  $\varepsilon > 0$  and a.s. finite random variable  $C(\omega) > 0$ . Then

$$\prod_{j=N}^i (1 - 2H_1\alpha_j)^{-1} = \prod_{j=N}^i e^{-\ln(1-2H_1\alpha_j)} = e^{-\sum_{j=N}^i \ln(1-2H_1\alpha_j)} \leq K'_1 i^{2H_1},$$

and a.s.

$$\sum_{i=1}^{\infty} \alpha_i \eta_{i+1} \prod_{j=1}^i (1 - 2\alpha_j H_1)^{-1} \leq K'_2 \sum_{i=1}^{\infty} \frac{1}{i} i^{2H_1} \frac{C(\omega)}{i^{\varepsilon+2H_1}} < \infty,$$

where  $K'_1$  and  $K'_2$  are some constants.

b) Let  $\alpha_i = (i \ln i)^{-1}$  and  $\eta_i \leq C(\omega)/i^{\varepsilon+2H_1}$  for some  $\varepsilon > 0$  and a.s. finite random variable  $C(\omega) > 0$ . Then

$$\prod_{j=N}^i (1 - 2H_1\alpha_j)^{-1} \leq K'_1 \ln^{2H_1} i,$$

and a.s.

$$\sum_{i=0}^{\infty} \alpha_i \eta_{i+1} \prod_{i=1}^i (1 - 2\alpha_i H_1)^{-1} \leq K'_2 \sum_{i=1}^{\infty} \frac{1}{i \ln i} \ln^{2H_1} i \frac{C(\omega)}{i^{\varepsilon+2H_1}} < \infty.$$

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